



## Sex-specific individual and joint associations of multiple environmental exposures with diabetes and obesity in the population-based German National Cohort (NAKO)

Fiona Niedermayer<sup>a,b,\*</sup> , Barbara Hoffmann<sup>c</sup> , Boya Zhang<sup>d,e</sup>, Jie Chen<sup>d,e</sup>, Jaime E. Hart<sup>d,e</sup> , Francine Laden<sup>d,e</sup> , Gabriele Bolte<sup>f,g</sup> , Tobia Lakes<sup>h,i</sup>, Tamara Schikowski<sup>j</sup> , Karin Halina Greiser<sup>k</sup>, Jeroen Staab<sup>l</sup>, Nikolaos Nikolaou<sup>b</sup> , Marco Dallavalle<sup>b</sup>, Matthias B. Schulze<sup>m,n</sup>, Wolfgang Lieb<sup>o</sup> , Cara Övermöhle<sup>o</sup>, Thaddäus Tönnies<sup>p</sup>, Verena Katzke<sup>k</sup> , Heiko Becher<sup>q</sup> , Beate Fischer<sup>r</sup>, Michael Leitzmann<sup>r</sup>, Klaus Berger<sup>s</sup>, Fatemeh Mayvaneh<sup>s</sup>, Thomas Keil<sup>t,u,v</sup> , Lilian Krist<sup>t</sup> , Carolina J. Klett-Tammen<sup>w</sup> , Jana-Kristin Heise<sup>w</sup> , Tobias Pischon<sup>x,y,z</sup>, Ilais Moreno Velásquez<sup>x</sup>, Børge Schmidt<sup>aa</sup>, Rajini Nagrani<sup>ab</sup>, Stefan Rach<sup>ab</sup>, Hermann Brenner<sup>ac</sup>, Bernd Holleccek<sup>ad</sup>, Volker Harth<sup>ae</sup>, Nadia Obi<sup>ae</sup> , Anna Köttgen<sup>af</sup>, Rafael Mikolajczyk<sup>ag</sup> , Claudia Meinke-Franze<sup>ah</sup>, Wolfgang Hoffmann<sup>ah</sup> , Alexandra Schneider<sup>b</sup> , Kathrin Wolf<sup>b</sup> , Annette Peters<sup>a,b,ai</sup>

<sup>a</sup> Chair of Epidemiology, IBE, Faculty of Medicine, Ludwig-Maximilians-Universität Munich, Munich, Germany

<sup>b</sup> Institute of Epidemiology, Helmholtz Zentrum München (GmbH) - German Research Center for Environmental Health, Neuherberg, Germany

<sup>c</sup> Institute of Occupational, Social and Environmental Medicine, Medical Faculty and University Hospital, Heinrich Heine University Düsseldorf, Düsseldorf, Germany

<sup>d</sup> Department of Environmental Health, Harvard T. H. Chan School of Public Health, Boston, MA, USA

<sup>e</sup> Channing Division of Network Medicine, Department of Medicine, Brigham and Women's Hospital and Harvard Medical School, Boston, MA, USA

<sup>f</sup> Department of Social Epidemiology, Institute of Public Health and Nursing Research, University of Bremen, Bremen, Germany

<sup>g</sup> Health Sciences Bremen, University of Bremen, Bremen, Germany

<sup>h</sup> Geography Department, Geoinformation Science Lab, Humboldt Universität zu Berlin, Berlin, Germany

<sup>i</sup> Integrative Research Institute on Transformations of Human Environment Systems (IRI THESys), Berlin, Germany

<sup>j</sup> IUF – Leibniz Research Institute for Environmental Medicine, Düsseldorf, Germany

<sup>k</sup> Division of Cancer Epidemiology (CO20), German Cancer Research Center (DKFZ), Heidelberg, Germany

<sup>l</sup> German Remote Sensing Data Center, German Aerospace Center (DLR), Weßling, Germany

<sup>m</sup> Department of Molecular Epidemiology, German Institute of Human Nutrition Potsdam-Rehbruecke, Nuthetal, Germany

<sup>n</sup> Institute of Nutritional Science, University of Potsdam, Nuthetal, Germany

<sup>o</sup> Institute of Epidemiology, Kiel University, Kiel, Germany

<sup>p</sup> Institute for Biometrics and Epidemiology, German Diabetes Center (DDZ), Leibniz Center for Diabetes Research at Heinrich Heine University, Düsseldorf, Germany

<sup>q</sup> Institute of Global Health, University Hospital Heidelberg, Heidelberg, Germany

<sup>r</sup> Department of Epidemiology and Preventive Medicine, University of Regensburg, Regensburg, Germany

<sup>s</sup> Institute of Epidemiology and Social Medicine, University of Münster, Münster, Germany

<sup>t</sup> Institute of Social Medicine, Epidemiology, and Health Economics, Charité - Universitätsmedizin Berlin, Berlin, Germany

<sup>u</sup> Institute of Clinical Epidemiology and Biometry, University of Würzburg, Würzburg, Germany

<sup>v</sup> State Institute of Health I, Bavarian Health and Food Safety Authority, Erlangen, Germany

<sup>w</sup> Department of Epidemiology, Helmholtz Centre for Infection Research (HZI), Braunschweig, Germany

<sup>x</sup> Max-Delbrück-Center for Molecular Medicine in the Helmholtz Association (MDC), Molecular Epidemiology Research Group, Berlin, Germany

<sup>y</sup> Max-Delbrück-Center for Molecular Medicine in the Helmholtz Association (MDC), Biobank Technology Platform, Berlin, Germany

<sup>z</sup> Charité - Universitätsmedizin Berlin, Corporate member of Freie Universität Berlin and Humboldt-Universität zu Berlin, Berlin, Germany

<sup>aa</sup> Institute for Medical Informatics, Biometry and Epidemiology, University Hospital of Essen, University of Duisburg-Essen, Essen, Germany

<sup>ab</sup> Department Epidemiological Methods and Etiological Research, Leibniz Institute for Prevention Research and Epidemiology - BIPS, Bremen, Germany

<sup>ac</sup> Cancer Prevention Graduate School, German Cancer Research Center (DKFZ), Heidelberg, Germany

<sup>ad</sup> NAKO Study Center Saarbrücken, Saarland Cancer Registry, Saarbrücken, Germany

\* Corresponding author. Ingolstädter Landstr. 1, 85764, Neuherberg, Germany.

E-mail address: [fiona.niedermayer@helmholtz-munich.de](mailto:fiona.niedermayer@helmholtz-munich.de) (F. Niedermayer).

<sup>ae</sup> Institute for Occupational and Maritime Medicine (ZfAM), University Medical Center Hamburg-Eppendorf, Hamburg, Germany

<sup>af</sup> Institute of Epidemiology and Prevention, Faculty of Medicine and Medical Center – University of Freiburg, Freiburg, Germany

<sup>ag</sup> Institute for Medical Epidemiology, Biometrics, and Informatics, Interdisciplinary Center for Health Sciences, Medical Faculty of the Martin Luther University Halle-

Wittenberg, Halle (Saale), Germany

<sup>ah</sup> Institute for Community Medicine, University Medicine Greifswald, Greifswald, Germany

<sup>ai</sup> German Center for Diabetes Research (DZD e.V.), Neuherberg, Germany

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## ABSTRACT

Recent studies have suggested a potential association of particulate matter (PM) and noise with diabetes and obesity, but studies examining other environmental exposures and their sex-specific and joint associations remain limited. Therefore, we investigated sex-specific individual and joint associations of annual exposure to multiple environmental factors with diabetes and obesity-related measures using cross-sectional data from the population-based multi-center German National Cohort (NAKO). Outcomes included self-reported diabetes mellitus, body mass index (BMI), obesity (BMI  $\geq 30$  kg/m<sup>2</sup>), and waist circumference. Annual mean residential exposures included air pollutants, air temperature, day-evening-night road traffic noise (L<sub>den</sub>) and surrounding greenness (normalized difference vegetation index (NDVI)). We used sex-stratified linear and logistic regression models to assess individual associations and quantile g-computation to assess joint associations. Among 174,955 adult participants (50.4% women), 5.6% reported a diabetes diagnosis and 20.9% were obese. An interquartile range increase in PM<sub>2.5</sub> and L<sub>den</sub> was consistently associated with diabetes and obesity-related measures (e.g., PM<sub>2.5</sub>-diabetes for men: odds ratio (OR) [95% confidence interval] = 1.12 [1.02; 1.22]; L<sub>den</sub>-BMI for women: 0.22 kg/m<sup>2</sup> [0.16; 0.27]). Greenness showed non-linear (inverted U-shaped) with all outcomes. An interquartile range increase in multiple exposures simultaneously was associated with higher odds of diabetes, obesity and higher obesity-related measures (e.g., mixture (PM<sub>2.5</sub>, L<sub>den</sub>, lack of NDVI)-diabetes: OR = 1.20 [1.09; 1.33] for men; mixture (PM<sub>2.5</sub>, L<sub>den</sub>, lack of NDVI)-BMI: 0.33 kg/m<sup>2</sup> [0.21; 0.44] for women). While longitudinal studies need to confirm these findings, the study highlights that reducing multiple adverse environmental exposures could be potential targets for the prevention of diabetes and obesity.

### Abbreviations

<b>BMI</b>	Body mass index
<b>CI</b>	Confidence Interval
<b>CVD</b>	Cardiovascular disease
<b>DAG</b>	Directed acyclic graph
<b>EIONET</b>	European Environment Information and Observation Network
<b>ELAPSE</b>	Effects of Low-Level Air Pollution: A Study in Europe
<b>END</b>	Environmental Noise Directive
<b>ISCED</b>	International Standard Classification of Education
<b>L<sub>den</sub></b>	Day-evening-night road traffic noise level
<b>NAKO</b>	German National Cohort
<b>NDVI</b>	Normalized Difference Vegetation Index
<b>NO<sub>2</sub></b>	Nitrogen dioxide
<b>OR</b>	Odds ratio
<b>PM</b>	Particulate matter
<b>T<sub>mean</sub></b>	Mean air temperature
<b>SES</b>	Socioeconomic status
<b>WHO</b>	World Health Organization

## 1. Introduction

With over 529 million people diagnosed with diabetes mellitus and 650 million affected by obesity (body mass index (BMI)  $\geq 30$  kg/m<sup>2</sup>) in early 2020s (Global Burden of Disease, 2021 Diabetes Collaborators, 2023; Sorensen et al., 2022a), metabolic diseases represent a serious public health challenge worldwide. Moreover, obesity constitutes a risk factor for a wide range of other non-communicable diseases, including cardiovascular disease (CVD), cancer, and type 2 diabetes (Bluher, 2019). The incidence of metabolic diseases continues to rise globally (Global Burden of Disease, 2021 Diabetes Collaborators, 2023; Non-Communicable Disease Risk Factor Collaboration, 2024), despite the

implementation of behavioral interventions designed to increase physical activity and to improve dietary habits, which are primary strategies in the prevention of metabolic diseases (Bluher, 2019). Low socioeconomic status (SES) has been identified as a crucial risk factor for diabetes and obesity (Schienkiewitz et al., 2022; Hoffmann et al., 2017), which often coincides with further socio-economic and environmental disparities due to more disadvantaged living conditions.

Environmental factors such as air pollution, traffic noise, ambient air temperature or lack of surrounding greenness are crucial health determinants (Global Burden of Disease Risk Factors Collaborators, 2024; Rojas-Rueda et al., 2021). In the latest update of the Global Burden of Disease Study in 2024, air pollution was identified as the most important risk factor attributable to deaths and disability-adjusted life years from diabetes, kidney disease, respiratory disease and CVD (Global Burden of Disease Risk Factors Collaborators, 2024). The World Health Organization (WHO) denoted traffic noise as the second most harmful environmental risk factor, evidencing levels above 53 dB(A) to adverse health effects. (WHO, 2018) As environmental risk factors can partly be controlled at the population level, for example by legally binding limit values for air pollution and traffic noise, or by considering re-vegetation of urban areas, these approaches could constitute powerful and essential public health interventions to prevent diabetes and obesity. Therefore, there is an urgent need to elucidate the potential impact of environmental factors on diabetes and obesity-related measures.

Growing evidence links high levels of air pollution, especially particulate matter (PM), and road traffic noise to diabetes prevalence and incidence (He et al., 2017; Yang et al., 2020a; Wang et al., 2020). However, the evidence on other air pollutants and environmental factors, including air temperature and greenness, are less clear and presently understudied (Valdes et al., 2019; De la Fuente et al., 2020). Furthermore, evidence on the association of environmental factors with obesity-related measures is inconclusive (An et al., 2018; Luo et al., 2020; Gui et al., 2022; Huang et al., 2020). For example, a meta-analysis from 2018 reported that 56% of the included studies did not find adverse associations of air pollution with obesity markers (An et al., 2018). High exposure to air pollutants and noise may contribute to diabetes and obesity by causing systemic inflammation, sleep disturbance, and constant release of stress hormones, while beneficial effects on mental

health and stress reduction are attributed to higher levels of greenness (Rajagopalan and Brook, 2012; Rajagopalan et al., 2024; Munzel et al., 2024; Fong et al., 2018).

In addition, studies have tended to focus on individual exposures, such as air pollution or greenness, rather than examining associations of multiple environmental exposures. This may not adequately capture the true health effects of these environmental factors, as they may confound and interact with each other (Klompaker et al., 2019). It also ignores the fact that these factors co-exist, especially in urban areas (Nieuwenhuijsen, 2020). Due to the common traffic source, urban areas are often characterized by adverse levels of air pollution and road traffic noise while surrounding greenness is lacking, exposing urban residents to multiple environmental risk factors at the same time. There are only a few previous studies that looked at joint associations of multiple environmental exposures and metabolic diseases (Klompaker et al., 2019; Sorensen et al., 2022b; Zhang et al., 2024). Although different methods were used, all studies indicated stronger adverse associations when exposures were considered jointly (Klompaker et al., 2019; Sorensen et al., 2022b; Zhang et al., 2024). As study regions in previous studies were often restricted to urban areas (Pyko et al., 2017; Altug et al., 2024; Yang et al., 2019), there is insufficient evidence on which of these associations can still be observed in more rural and remote areas. Based on our previous analysis using data from another smaller German cohort (KORA-Fit), the associations of air pollutants and greenness with obesity-related measures varied by the degree of urbanization and sex (Niedermayer et al., 2024). As the previous study was limited to a smaller sample size ( $n = 3034$ ), we aimed to confirm these observations in a larger population. Moreover, we observed interactions by sex, indicating potential sex-specific susceptibility to environmental factors. As biological differences in metabolic regulation between men and women are well established (Kautzky-Willer et al., 2023; Chew et al., 2023), studies explicitly examining sex-specific environmental associations with metabolic related outcomes remain scarce.

Earlier German cohort studies contributed substantially to understanding environmental determinants of health, primarily through single exposure analyses (Peters et al., 1997; Schikowski et al., 2010; Hoffmann et al., 2007). However, these studies lacked spatial variability due to their restricted study regions, and their comparatively small sample sizes reduced statistical power, particularly for subgroup analyses (Wolf et al., 2025). However, the German National Cohort (NAKO), which was initiated in 2014 and comprises more than 200,000 extensively characterized participants, provides a much stronger basis for comprehensive analyses by integrating individual-level data with a broad range of residential environmental exposures, including air pollution, road traffic noise, meteorology, and indicators of the built environment. Therefore, we investigated the sex-specific associations of various environmental exposures in single-exposure and multi-exposure models taking into account urbanicity. We analyzed cross-sectional data from a multi-center, population-based cohort study and addressed the following research questions:

1. How are higher annual levels of environmental exposures associated with diabetes and obesity-related measures in men and women?
2. Are there joint associations of multiple environmental exposures with diabetes and obesity-related measures in men and women?
3. Do these associations vary in urban, suburban and rural areas?

## 2. Methods

### 2.1. Study design and population

The German National Cohort (NAKO Gesundheitsstudie) is a nationwide, population-based, multi-center cohort that aims to investigate a range of chronic diseases, their risk factors, and etiology. A detailed description of the NAKO, including information on study design and participant recruitment, can be found elsewhere (Peters et al., 2022;

Rach et al., 2025). Briefly, between 2014 and 2019, 205,415 participants aged 19-75 years underwent a baseline examination at one of 18 NAKO study centers within Germany. Participants were recruited in 16 different study center regions in Germany from the general population based on a sex and age-stratified designs (Rach et al., 2025), that comprised urban and rural living environments. Standardized, comprehensive interviews, medical and physical examinations were carried out by trained staff at the study centers. For the present analyses, we used cross-sectional data from 204,687 eligible participants who attended the baseline examination. The NAKO was approved by the local ethics committees and all participants gave written informed consent before study enrollment.

### 2.2. Outcomes

We used diabetes mellitus, obesity, BMI and waist circumference as outcome variables. The definition of diabetes mellitus was based on self-report. Participants were asked "Have you ever been diagnosed with diabetes mellitus by a physician?" during standardized, computer-assisted interviews at baseline examination (Jaeschke et al., 2020). BMI was calculated from standardized measurements of height and weight by dividing weight by the square of height ( $\text{kg}/\text{m}^2$ ) (Fischer et al., 2020). If height and/or weight measurements were missing (<1%), values were filled in with self-reported height and weight information. Obesity was defined as  $\text{BMI} \geq 30 \text{ kg}/\text{m}^2$  (WHO, 2000). Waist circumference was measured according to the WHO guideline using a measuring tape (Seca) placed midway between the lowest rib and the top of the iliac crest (WHO, 2008).

### 2.3. Exposures

We used the following environmental exposures: annual means of ambient air pollution, ambient air temperature, road traffic noise and greenness (Fig. S1). All environmental factors were assigned to the geocoded residential baseline addresses of participants at the time of the baseline examination. Detailed information on the exposure assessment, linkage, harmonization and data sources used can be found in Wolf et al. (2025).

Briefly, we used air pollution estimates from the ELAPSE (Effects of Low-Level Air Pollution: A Study in Europe) project, which provides annual levels of nitrogen dioxide ( $\text{NO}_2$ ), particulate matter  $<2.5 \mu\text{m}$  in diameter ( $\text{PM}_{2.5}$ ), and  $\text{PM}_{2.5}$  absorbance ( $\text{PM}_{2.5} \text{ abs}$ ), a proxy for black carbon, for Western Europe for the year 2010 (Wolf et al., 2025; de Hoogh et al., 2018). Land use regression models were developed to predict air pollution concentrations, incorporating measurements from ground-level monitoring stations, satellite observations, estimates from chemistry transport models, and further spatial predictors (de Hoogh et al., 2018). These models were then applied to compile country-wide maps with a resolution of  $100\text{m} \times 100\text{m}$  (de Hoogh et al., 2018).

Data on annual mean, minimum and maximum air temperature were available for the years 2000-2020 with a resolution of  $1 \text{ km} \times 1 \text{ km}$ . Air temperature measurements from ground-based measurement stations, satellite-derived land surface temperature and spatial factors were combined in a three-stage spatiotemporal model as described by Niko-laou et al. (2023). We assigned each participant the annual mean temperature ( $T_{\text{mean}}$ ) in the examination year.

Road traffic noise data at a scale of  $10\text{m} \times 10\text{m}$  for the year 2017 was available from the European Environment Information and Observation Network (EIONET) (Staab et al., 2025; European Commission, 2002). The maps provide day-evening-night levels ( $L_{\text{den}}$ ) categorized into 5 groups: (55: 55-59 dB(A), 60: 60-64 dB(A), 65: 65-69 dB(A), 70: 70-74 dB(A), 75:  $\geq 75$  dB(A)). Participants with missing data or living in urbanized areas not covered by the Environmental Noise Directive (END) obligation 2002/49/EG Article 3 section k<sup>47</sup> were set to 40 dB(A) defining a lower detection limit. We calculated mean values of grid cells within a 100m buffer size around residential addresses to obtain continuous

area-weighted noise levels (Wolf et al., 2025).

Lastly, we used the normalized difference vegetation index (NDVI) as indicator of surrounding greenness. Monthly data on a 1 km\*1 km grid were gathered from the NASA Terra Moderate Resolution Imaging Spectroradiometer (MODIS), available on a 1 km\*1 km grid, for the years 2014 to 2019 (Wolf et al., 2025). Briefly, the NDVI (= reflected radiation in the visible red minus in the near infrared divided by the sum of the two) takes values from -1 to 1. Prior to analysis, water pixels (values < 0) were excluded by masking the data with a mask layer, so values close to 0 indicate grey, less green areas, while values close to 1 indicate dense vegetation (Wolf et al., 2025). We assigned each participant the annual NDVI of the respective grid cell averaged over the vegetation period (March to October) of the examination year to reflect the whole vegetation period of this year.

#### 2.4. Covariables

Information on sociodemographic, lifestyle factors, and SES were assessed by standardized interviews, questionnaires, and physical examinations at baseline. Alcohol consumption is provided as gram per day. Participants were categorized as never, former and current smoker based on their self-reported smoking behavior. Physical activity was assessed by the Global Physical Activity Questionnaire, evaluated according to the WHO analysis guide (Leitzmann et al., 2020). We winorized implausible values to a maximum of 16 h a day per week (6720 min), and categorized into quintiles. We considered educational levels rather than income to reflect individual SES, as there were fewer missing values in this variable. Educational levels were classified by combining information on school and vocational training according to the International Standard Classification of Education (ISCED) 2011 (UNESCO Institute for Statistics, 2012) and we grouped participants still in education or training into a separate category.

Further geocoded variables assigned to the baseline residential addresses included neighborhood SES, population density and degree of urbanization. Neighborhood SES was represented by the unemployment rate at the district level for 2014 from the Federal Employment Agency (Bundesagentur für Arbeit). Population density, given as inhabitants per square meters within a 1 km buffer of participants' residences, was available for the year 2018 from a private company (WiGeoGis GmbH). Information on the degree of urbanization was downloaded from EUROSTAT (statistical office of the European Union) for the year 2020 (population census data 2018), which provides a categorization of German municipalities into urban, suburban, and rural areas based on the proportion of densely populated grid cells in each municipality (European Union, 2018). We used this information to define the degree of urbanization of participants' municipalities of residence, which we considered as potential effect modifier.

#### 2.5. Statistical analysis

We performed Spearman's rank correlation to assess correlations between environmental exposures and population density.

As our study aimed to report sex-specific associations, we applied linear and logistic regression models stratified by sex, adjusting for age, study center, lifestyle factors, SES and population density for our main analysis. Our basic model included outcome-related risk factors, such as age, smoking, physical activity, alcohol consumption, which explained some heterogeneity between our participants. We stepwise added the following confounders identified by directed acyclic graphs (DAG) (Fig. S2): study center, individual SES, neighborhood SES and population density. As neighborhood and study center region were relatively coarse, we hypothesized that these indicators may not fully capture socioeconomic heterogeneity within urban areas. Therefore, we included population density which may serve as a proxy for built environmental features. Degree of urbanization, however, was not included as a confounder. Instead, it was incorporated exclusively as an effect

modifier to examine whether associations differed across levels of urbanicity. In addition, we examined the shape of the exposure-response functions between each exposure and outcome to identify non-linear associations. We fitted cubic splines for the exposure using the *gam* function of the R package *mgcv* (Wood, 2011). If there was a visual detection of non-linearity in the main exposure ranges (10<sup>th</sup> – 90<sup>th</sup> percentile), we present plots showing the non-linear exposure-response curve. In order to get a proximation of the association between each exposure range and the turning points, we applied piecewise linear and logistic regression models, which is described in detail in the supplement (supplementary methods S1). Odds ratios (OR) and absolute changes derived from the sex-stratified regression models are reported per interquartile range (IQR) increase in environmental exposure with 95% confidence intervals (95% CI) in the main text. We repeated the single-exposure analysis in the total sample and results are provided in the supplementary material.

Furthermore, we exploratory assessed joint associations of an environmental exposure mixture with diabetes and obesity-related measures in sex-stratified multi-exposure models using quantile g-computation. This method has been described by Keil et al. (2020) and is suitable for the environmental exposure setting (Zhang et al., 2024; Chen et al., 2024; Stevens et al., 2023). Briefly, exposures of interest are standardized and categorized into a predefined increment, such as quartiles. In a supervised approach, exposures are ranked according to their importance and contribution to the outcome of interest and a weighted sum index is created. The outcome is regressed on this single index so that the joint association can be interpreted as the change in outcome if all exposures were simultaneously increased by one pre-defined increment. Since our noise variable could not be divided into quartiles due to tied values in the lowest quartile, we turned off the quantization and standardized all environmental exposures by dividing them by their IQR. Consequently, the joint associations from the quantile g-computation approach can be interpreted as the change in outcome if all exposures were simultaneously increased by one IQR. In addition, the weights indicate the proportion to which an exposure contributes to the positive or negative joint association. We implemented quantile g-computation using the R package *qgcomp* (Keil et al., 2020) and considered two different exposure mixture sets: (1) including exposures that showed significant associations in single-exposure models, (2) including all exposures. The inverse of NDVI was used because it was negatively correlated with air pollution, air temperature and road traffic noise and because we expected negative associations with diabetes and obesity-related measures. In addition, we ran two sensitivity analyses for each exposure mixture set. (1) We dropped the exposure that displayed non-linear exposure-response functions from the exposure mixture set. (2) We used natural cubic splines for exposures that displayed non-linear exposure-response functions in single-exposure regression models. However, we need to note that for these exposure mixtures including splines, the method cannot derive weights attributed to each exposure. We also ran two-exposure models, mutually adjusting each environmental factor for another, to assess the independence of associations. Therefore, we paired exposures with Spearman correlation coefficients <0.7 to avoid multicollinearity.

We further assessed effect modification by adding an interaction term between exposure and degree of urbanization (urban/suburban/rural). In addition, we analyzed regression models stratified by study center region to identify between-center heterogeneity among associations.

We assessed the robustness of our results by performing several sensitivity analyses. Firstly, we used different adjustment sets: (1) we additionally adjusted for income and partnership, (2) we did not adjust for physical activity, as this could be a mediator in the association of NDVI with diabetes and obesity (De la Fuente et al., 2020), (3) we included BMI as a potential mediator in single-exposure models with diabetes as the outcome, in order to reflect the direct rather than total association between exposure and diabetes. Secondly, we excluded

participants who had self-reported a cancer or diabetes diagnosis when considering obesity-related outcomes. Thirdly, we excluded participants who reported living at the current residential address for less than five or ten years. Finally, we applied different definitions of diabetes on the associations: (1) we excluded all women who reported a diagnosis of diabetes onset during pregnancy (gestational diabetes) ( $n = 1157$ ); (2) we excluded cases with probable type 1 diabetes (defined as age of diagnosis  $\leq 30$  years according to the DIAB-CORE (Schunk et al., 2012)), (3) we excluded cases which had only reported a diabetes diagnosis, but not the use of anti-diabetic medication.

All analyses were done with RStudio Version 4.3.1, and p-values  $< 0.05$  indicated statistical significance. All analyses were performed separately for each outcome.

### 3. Results

#### 3.1. Sample characteristics

After excluding all individuals with missing data on any outcome, exposure or main covariates, our final analytic sample consisted of 174,955 participants (Fig. S3), of whom 50.4% were women, the mean age was 49.5 years, and more than 70% lived in urban areas (Table 1). In the present sample, 6.2% of the men and 5.0% of women reported a diabetes diagnosis, with a mean age at diagnosis of 50 and 45 years, respectively. For men, the mean BMI and waist circumference were 27.2  $\text{kg}/\text{m}^2$  and 96.5 cm, respectively, and 22% were obese. For women, the mean BMI and waist circumference were 26.0  $\text{kg}/\text{m}^2$  and 85.6 cm, respectively, and 19.8%, were obese. Women reported lower alcohol consumption, were less likely to be ex- or current smokers had lower levels of physical activity, were less likely to have a university degree and lived in more densely populated areas compared to men (Table 1). A comparison of the characteristics of the excluded and included participants can be seen in Table S1.

Environmental exposure levels were similar for men and women but differed according to the study region. Mean levels of  $\text{NO}_2$ ,  $\text{PM}_{2.5}$  and  $\text{PM}_{2.5}$  abs were 26.8  $\mu\text{g}/\text{m}^3$ , 17.3  $\mu\text{g}/\text{m}^3$ , and  $1.6 \cdot 10^{-5} \text{ m}^{-1}$ , respectively (Table S2). Mean air temperature was 10.9 °C, and the participants were exposed to a mean road traffic noise level of 44.6 dB(A) and a mean surrounding greenness cover of 0.5. Correlations were positive between air pollution, air temperature, and noise ( $r = 0.35$  to 0.84), NDVI was negatively correlated with all other environmental exposures ( $r = -0.59$  to  $-0.34$ ) (Table S2).  $L_{\text{den}}$  showed the weakest correlation with  $\text{PM}_{2.5}$  among the air pollutants ( $r = 0.35$ , compared with  $r = 0.61$  for  $\text{NO}_2$  and  $r = 0.56$  for  $\text{PM}_{2.5}$  abs) and was only modestly negatively correlated with NDVI ( $r = -0.34$ ). Population density was positively correlated with air pollution, air temperature, and noise ( $r = 0.42$  to 0.72) and negatively with NDVI ( $r = -0.64$ ).

#### 3.2. Single-exposure models

Exposure-response functions showed linear forms in the main exposure ranges (10<sup>th</sup> – 90<sup>th</sup> percentile) for air pollution, temperature, and noise (Figs. S4–S7) with diabetes and obesity-related measures, therefore we show linear effect estimates (Tables 2–3). Estimates changed strongest when population density were included in the model, followed by study center region, and education (Tables S3–S4).

For men and women, an IQR increase in annual  $\text{PM}_{2.5}$  was associated with higher odds of diabetes (e.g., men: OR = 1.12 [95% CI: 1.02; 1.22], women: OR = 1.11 [1.01; 1.22]; per 2.9  $\mu\text{g}/\text{m}^3$  increase). An association of higher  $\text{NO}_2$  and  $\text{PM}_{2.5}$  abs was only found for diabetes in men (OR ranged from 1.08 to 1.10; Table 2). For women only,  $T_{\text{mean}}$  was associated with higher odds of diabetes, while an increase of 8.3 dB(A) in annual  $L_{\text{den}}$  was associated with 1.08-fold and 1.05-fold higher odds of diabetes in men and women, respectively. With regard to the obesity-related measures, an IQR increase in annual  $\text{PM}_{2.5}$  was associated with higher obesity-related measures (Tables 2 and 3). For example, higher

**Table 1**

Participants' characteristics in the final analytical sample of the German National Cohort (NAKO), overall and stratified by sex.

	Overall (n = 174,955)	Men (n = 86,710)	Women (n = 88,245)
<b>Sex, female n (%)</b>	88,245 (50.4)	-	88,245 (100.0)
<b>Age (years), mean (SD)</b>	49.5 (12.8)	49.6 (12.8)	49.5 (12.7)
<b>Diagnosis of diabetes, yes n (%)</b>	9,747 (5.6)	5,352 (6.2)	4,395 (5.0)
<b>Age at diabetes diagnosis, mean (SD)</b>	47.6 (13.1)	49.9 (12.0)	44.9 (13.8)
<b>Antidiabetic medication, n (%)</b>	6,695 (3.8)	4,221 (4.9)	2,474 (2.8)
<b>Obesity (<math>\geq 30 \text{ kg}/\text{m}^2</math>) yes n (%)</b>	36,579 (20.9)	19,096 (22.0)	17,483 (19.8)
<b>Body Mass Index (<math>\text{kg}/\text{m}^2</math>), mean (SD)</b>	26.6 (5.0)	27.19 (4.5)	26.0 (5.5)
<b>Waist circumference (cm), mean (SD)</b>	91.0 (14.3)	96.5 (12.9)	85.6 (13.5)
<b>Physical activity (min/week), mean (SD)</b>	1,440 (1,583)	1,466 (1,621)	1,415 (1,544)
<b>Quintiles of physical activity, n (%)</b>			
Q1	30,679 (17.5)	15,560 (17.9)	15,119 (17.1)
Q2	35,860 (20.5)	17,915 (20.7)	17,945 (20.3)
Q3	36,717 (21.0)	17,790 (20.5)	18,927 (21.4)
Q4	36,009 (20.6)	17,820 (20.6)	18,189 (20.6)
Q5	35,690 (20.4)	17,625 (20.3)	18,065 (20.5)
<b>Alcohol consumption (g/day), mean (SD)</b>	10.6 (16.7)	14.08 (19.4)	7.1 (12.6)
<b>Smoking behavior, n (%)</b>			
Never-smoker	81,151 (46.4)	36,374 (41.9)	44,777 (50.7)
Ex-smoker	56,988 (32.6)	30,547 (35.2)	26,441 (30.0)
Current smoker	36,816 (21.0)	19,789 (22.8)	17,027 (19.3)
<b>Education (ISCED 2011), n (%)</b>			
Primary	687 (0.4)	338 (0.4)	349 (0.4)
Lower secondary	3,527 (2.0)	1,274 (1.5)	2,253 (2.6)
Upper secondary	55,430 (31.7)	25,185 (29.0)	30,245 (34.3)
Post-secondary non-tertiary	15,700 (9.0)	6426 (7.4)	9274 (10.5)
Bachelor's	55,324 (31.6)	30,492 (35.2)	24,832 (28.1)
Master's	32,657 (18.7)	16,275 (18.8)	16,382 (18.6)
Doctoral	7,695 (4.4)	4,750 (5.5)	2,945 (3.3)
Still in education or training	3,935 (2.2)	1,970 (2.3)	1,965 (2.2)
<b>Income (Euros), mean (SD)</b>	3,721 (2,493)	3,975 (2,667)	3,466 (2,276)
<b>Partnership, n (%)</b>			
Single	32,807 (18.8)	14,051 (16.2)	18,756 (21.3)
Living with partner	124,039 (71.0)	63,983 (73.9)	60,056 (68.1)
Partner but living separately	17,894 (10.2)	8576 (9.9)	9318 (10.6)
<b>Unemployment rate at district level (%), mean (SD)</b>	8.62 (3.21)	8.60 (3.22)	8.63 (3.21)
<b>Population density (<math>\text{n}/\text{m}^2</math>), 1 km buffer, mean (SD)</b>	4652 (4363)	4628 (4350)	4675 (4375)
<b>Degree of urbanization at municipality level, n (%)</b>			
Urban	124,979 (71.4)	62,004 (71.5)	62,975 (71.4)
Suburban	27,923 (16.0)	13,771 (15.9)	14,152 (16.0)
Rural	22,045 (12.6)	10,933 (12.6)	11,112 (12.6)

\*Q1:  $\leq 210$  min/wk for men,  $\leq 200$  min/wk for women, Q2:  $>210 - \leq 540$  min/wk for men,  $>200 - \leq 530$  min/wk for women, Q3:  $>540 - \leq 1130$  min/wk for men,  $>530 - \leq 1110$  min/wk for women, Q4:  $>1130 - \leq 2490$  min/wk for men,  $>1110 - \leq 2370$  min/wk for women, Q5:  $>2490 - 6720$  min/wk for men,  $>2370 - 6720$  for women. Abbreviations: ISCED = International Standard Classification of Education, N = number, SD = standard deviation.

**Table 2**  
Associations of environmental exposures with existing diabetes and obesity from single-exposure logistic regression models in the German National Cohort (NAKO).

Exposure	IQR	Self-reported diagnosis of diabetes		Obesity (BMI ≥30 kg/m <sup>2</sup> )	
		Men (n = 86,710)	Women (n = 88,245)	Men (n = 86,710)	Women (n = 88,245)
		OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)
NO <sub>2</sub> [µg/m <sup>3</sup> ]	10.6	1.10 (1.02; 1.19)	1.04 (0.96; 1.13)	1.05 (1.01; 1.10)	1.05 (1.01; 1.10)
PM <sub>2.5</sub> [µg/m <sup>3</sup> ]	2.9	1.12 (1.02; 1.22)	1.11 (1.01; 1.22)	1.07 (1.02; 1.13)	1.10 (1.04; 1.16)
PM <sub>2.5</sub> abs [10 <sup>-5</sup> m <sup>-1</sup> ]	0.5	1.08 (1.01; 1.15)	1.03 (0.96; 1.11)	1.01 (0.97; 1.05)	0.98 (0.95; 1.02)
T <sub>mean</sub> [°C]	1.1	0.96 (0.90; 1.04)	1.10 (1.02; 1.19)	0.98 (0.94; 1.02)	1.00 (0.96; 1.04)
L <sub>den</sub> (100m) [dB(A)]	8.3	1.08 (1.03; 1.13)	1.05 (1.00; 1.10)	1.07 (1.04; 1.09)	1.08 (1.05; 1.11)

Legend: All models were adjusted for age, study center region, physical activity, alcohol consumption, smoking behavior, education, unemployment rate at district level and population density. ORs and confidence intervals are given per IQR increase in exposure. Diabetes was reported in n = 5352 men and n = 4395 women, and n = 19,096 men and n = 17,483 women were obese. Abbreviation: BMI = Body Mass Index, CI = confidence interval, IQR = interquartile range; L<sub>den</sub> = day-evening-night noise level, NO<sub>2</sub> = nitrogen dioxide, OR = odds ratio, PM<sub>2.5</sub> = particulate matter with diameter <2.5 µm; PM<sub>2.5</sub>abs = PM<sub>2.5</sub> absorbance, T<sub>mean</sub> = annual mean temperature.

**Table 3**  
Associations of environmental exposures with BMI and waist circumference from single-exposure linear regression models in the German National Cohort (NAKO).

Exposure	IQR	BMI		Waist circumference	
		Men (n = 86,710)	Women (n = 88,245)	Men (n = 86,710)	Women (n = 88,245)
		β (95% CI)	β (95% CI)	β (95% CI)	β (95% CI)
NO <sub>2</sub> [µg/m <sup>3</sup> ]	10.6	0.06 (-0.02; 0.13)	0.13 (0.04; 0.22)	0.24 (0.04; 0.44)	0.25 (0.04; 0.47)
PM <sub>2.5</sub> [µg/m <sup>3</sup> ]	2.9	0.11 (0.02; 0.20)	0.18 (0.07; 0.29)	0.40 (0.16; 0.64)	0.48 (0.21; 0.74)
PM <sub>2.5</sub> abs [10 <sup>-5</sup> m <sup>-1</sup> ]	0.5	-0.02 (-0.09; 0.04)	-0.01 (-0.08; 0.07)	0.01 (-0.16; 0.18)	-0.05 (-0.24; 0.14)
T <sub>mean</sub> [°C]	1.1	-0.06 (-0.12; 0.01)	-0.07 (-0.15; 0.01)	-0.25 (-0.43; -0.07)	-0.28 (-0.48; -0.08)
L <sub>den</sub> (100m) [dB(A)]	8.3	0.11 (0.06; 0.15)	0.22 (0.16; 0.27)	0.26 (0.15; 0.38)	0.46 (0.34; 0.59)

Legend: All models were adjusted for age, study center region, physical activity, alcohol consumption, smoking behavior, education, unemployment rate at district level and population density. Betas and confidence intervals are given as IQR increase in exposure. Abbreviation: BMI = Body Mass Index, CI = confidence interval, IQR = interquartile range, L<sub>den</sub> = day-evening-night noise level, NO<sub>2</sub> = nitrogen dioxide, PM<sub>2.5</sub> = particulate matter with diameter <2.5 µm, PM<sub>2.5</sub>abs = PM<sub>2.5</sub> absorbance, T<sub>mean</sub> = annual mean temperature.

annual PM<sub>2.5</sub> was associated with 0.11 kg/m<sup>2</sup> and 0.18 kg/m<sup>2</sup> higher BMI and 0.40 cm and 0.48 cm higher waist circumference in men and women, respectively. Except for T<sub>mean</sub> and waist circumference, the associations were close to unity for NO<sub>2</sub>, PM<sub>2.5</sub>abs and T<sub>mean</sub> for men and women. Higher annual L<sub>den</sub> was associated with higher obesity-related measures for both men and women (e.g., men: 0.11 kg/m<sup>2</sup> [0.06; 0.15], 0.26 cm [0.15; 0.38]; women: 0.22 kg/m<sup>2</sup> [0.16; 0.27], 0.46 cm [0.34; 0.59]; per 8.3 dB(A) increase). Effect estimates for the total sample are provided in Table S5.

A deviation from linearity was observed for NDVI with all outcomes, suggesting an inverted U-shaped function and a turning point around the median (0.55) (Fig. 1). Thus, indicating that low NDVI (below the median) and high NDVI (above the median) were associated with lower odds of diabetes, obesity, and lower BMI and waist circumference in men and women (Fig. 1). The inverted u-shape was independent of the adjustment model (data not shown) and also observed in the total sample (Fig. S8).

### 3.3. Joint associations of environmental exposure mixtures

For diabetes, we observed positive joint associations for men and women (Table S6). Considering an IQR increase in annual PM<sub>2.5</sub>, L<sub>den</sub> and an IQR decrease in NDVI simultaneously, these exposures were associated with ORs of 1.20 [1.09; 1.33] and 1.13 [1.01; 1.27] for diabetes in men and women, respectively (Fig. 2). For men, the positive joint association was equally driven by all three exposures. In women, it was driven by PM<sub>2.5</sub> (70% of the positive weight attributed) (Fig. 2). Although the magnitudes are not directly comparable, ORs from quantile g-computation indicated joint associations of similar or slightly larger magnitude than the single-exposure associations (e.g., ORs in single-exposure models ranged from: 1.08 - 1.12 per IQR increase). When all six environmental exposures were considered, the joint association was attenuated but remained because T<sub>mean</sub>, NO<sub>2</sub> and PM<sub>2.5</sub>abs contributed to the negative side, especially for men (Table S6; Fig. S9). Multi-exposure sets that excluded NDVI also showed similar, albeit slightly weaker, ORs for exposure mixture effects (Table S6). As NDVI displayed non-linear associations in the single-exposure models, we used natural cubic splines for NDVI in the two different mixture sets. However, this resulted in large confidence intervals and changes in the estimates (Table S6).

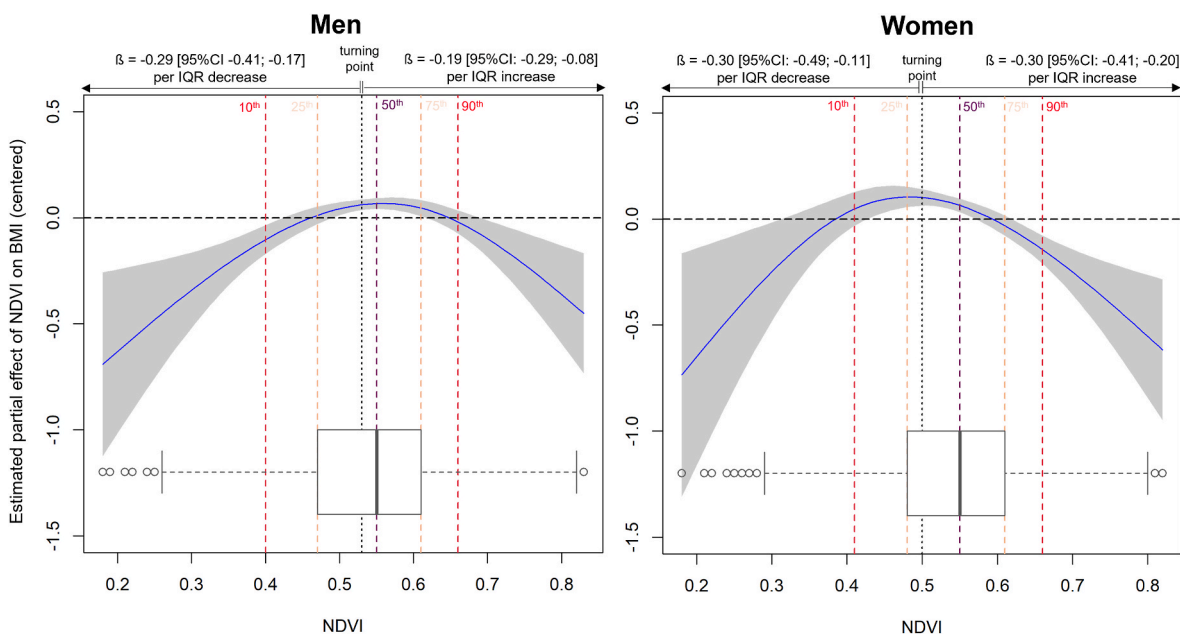
For BMI, a IQR increase in annual PM<sub>2.5</sub>, L<sub>den</sub> and lack of NDVI was jointly associated with 0.12 kg/m<sup>2</sup> [0.03; 0.21] and 0.33 kg/m<sup>2</sup> [0.21; 0.44] higher BMI in men and women, respectively (Fig. 2). For men and women, this joint association was mainly driven by L<sub>den</sub>, which accounted for more than 50% of the total positive association, followed by PM<sub>2.5</sub>. Similarly, joint ORs of obesity and increases in BMI and waist circumference associated with an IQR increase across the exposure mixtures had a similar or slightly larger magnitude than those from the single-exposure models. When all six environmental exposures were considered, the joint association was attenuated but remained except for BMI in men (Table S6, Fig. S9). Excluding NDVI<sub>inverse</sub> from the exposure mixture sets did not markedly alter the joint associations (Table S6). Using splines for NDVI, we observed changes in the estimates and large CI that lead to null findings (Table S6). Similar results were observed for waist circumference and obesity (Table S6).

Two-exposure models confirmed the main contribution of L<sub>den</sub>, as additional adjustments for other environmental exposures did not change the associations of L<sub>den</sub> with diabetes or obesity-related measures (Fig. S10). In contrast, associations of air pollutants attenuated after adjustment for L<sub>den</sub> (Fig. S10).

### 3.4. Exploratory analysis

#### 3.4.1. Effect modification by degree of urbanization

Baseline characteristics stratified by sex and degree of urbanization can be found in Table S7. Briefly, we observed a higher frequency of diabetes and obesity in rural areas than in urban areas among men and women. For diabetes, interactions between degree of urbanization and exposures were observed for men (Fig. S11). In suburban and rural areas, higher annual air pollution was positively associated with diabetes in men. In addition, higher annual L<sub>den</sub> in urban areas was associated with 1.05-fold and 1.06-fold higher odds of diabetes for men and women. For obesity-related measures, interactions between degree of urbanization and exposures were observed for men and women (e.g., BMI in Fig. 3). In urban and suburban areas, an IQR increase in annual



**Fig. 1.** Non-linear associations between NDVI and BMI in  $n = 86,710$  men and  $n = 88,245$  women from the German National Cohort (NAKO). Legend: Boxplots present distribution of exposure, vertical lines indicate NDVI percentiles (red, yellow, purple) and turning point (black). Models were fitted by cubic splines for the exposure and were additionally adjusted for age, study center region, lifestyle factors, education, unemployment rate at district level and population density. The betas and 95%-CI are given per IQR of 0.14 decrease in NDVI below the turning point ( $\leq 0.53$  for men,  $\leq 0.50$  for women) and per IQR of 0.14 increase in NDVI above the turning point ( $> 0.53$  for men,  $> 0.50$  for women) derived from piecewise linear regression models. Abbreviations: BMI = Body Mass Index, CI = confidence interval, IQR = interquartile range, NDVI = Normalized difference vegetation index. (For interpretation of the references to colour in this figure legend, the reader is referred to the Web version of this article.)

air pollutants and  $L_{den}$  were associated with higher BMI, waist circumference and obesity in men and women. In rural areas, we found no association of environmental factors with obesity-related measures except for  $NO_2$  in women. Non-linear associations between NDVI and obesity-related measures were driven by urban areas (e.g., BMI in Fig. S12).

### 3.4.2. Study center region-specific findings

The number of participants per study center region can be found in the Table S8. Radar plots showed different distribution patterns of environmental factors, BMI and population density across the study center regions (Fig. S13). Study center regions with a higher proportion of rural areas and lower population density had on average higher BMI and NDVI, but lower levels of air pollution,  $T_{mean}$  and  $L_{den}$  (e.g., Augsburg, Neubrandenburg, Saarbrücken, Regensburg). Study center regions with a large proportion of urban areas and high population density were characterized by lower mean BMI and NDVI, but higher air pollution and  $L_{den}$  (e.g., Berlin, Düsseldorf, Hannover, Mannheim). We observed a large heterogeneity in the associations when stratified by study center region (Figs. S14–15).

### 3.5. Sensitivity analyses

In sensitivity analyses, associations were robust. Using different confounder adjustment sets, consistent associations between adverse  $PM_{2.5}$  levels and  $L_{den}$  were observed for all outcomes (Tables S9–10). The associations with obesity outcomes did not change when we excluded existing diabetes and cancer cases (Table S11). When restricting the sample to those living at least five or 10 years at their residences, associations were robust (Tables S12–S13). For women, the associations were slightly stronger after excluding cases with gestational diabetes or cases which had only reported a diabetes diagnosis, but not the use of anti-diabetic medication (Table S14). For men, the associations slightly attenuated after excluding probable type 1 diabetes cases (Table S14).

## 4. Discussion

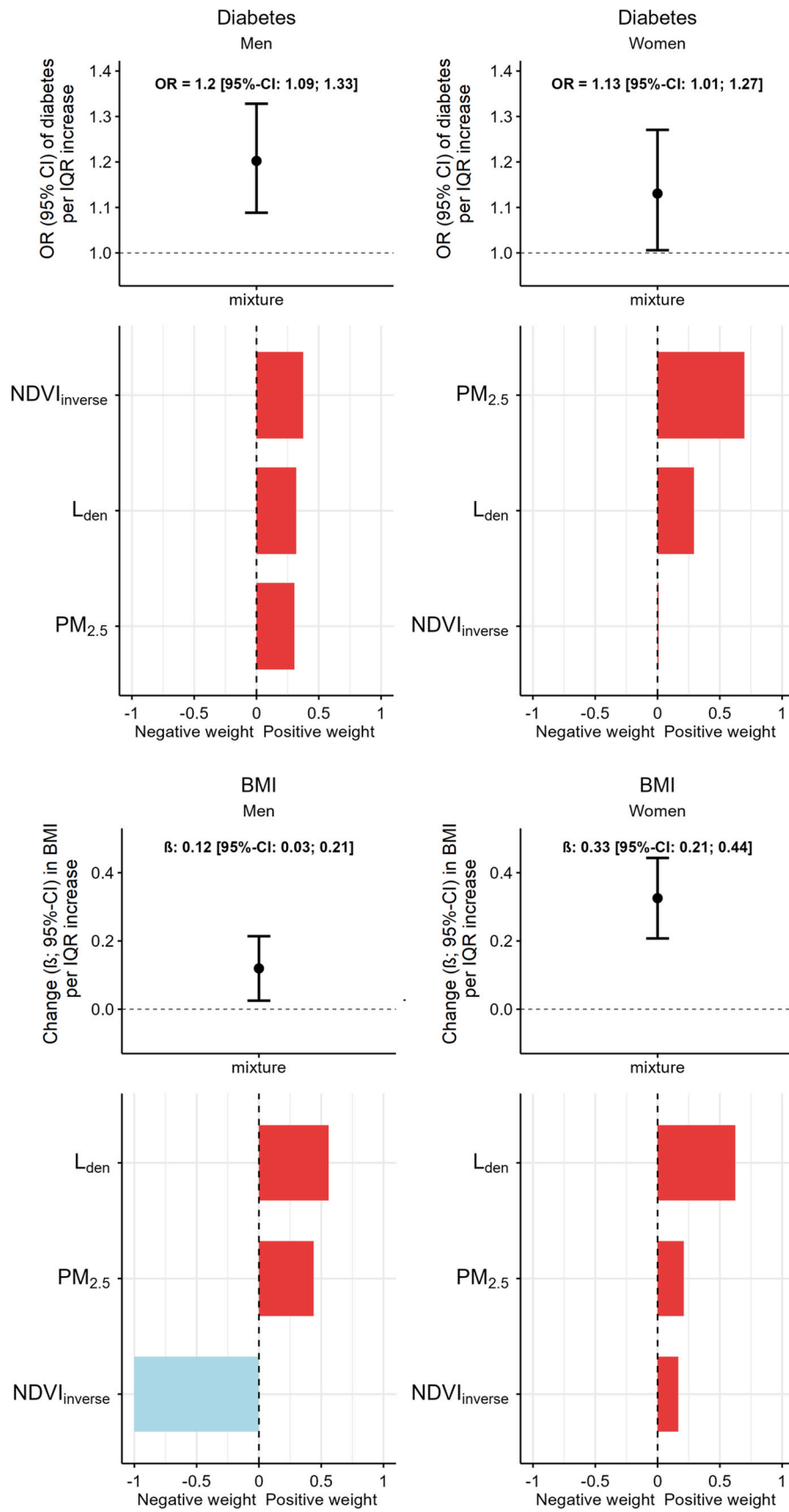
### 4.1. Summary and key points

In the current analyses, we assessed the associations of multiple environmental exposures with existing diabetes, obesity, BMI, and waist circumference using cross-sectional data from 174,955 participants of the multi-center NAKO study. We found consistent associations linking  $PM_{2.5}$  and road traffic noise to diabetes and obesity-related measures, especially in urban and suburban areas. NDVI showed an inverted U-shaped exposure-response with diabetes and obesity-related measures. Environmental exposure mixtures, when considering an IQR increase across all exposures or only  $PM_{2.5}$ , road traffic noise and lack of greenness, were associated with higher joint odds of diabetes, obesity and higher BMI and waist circumference, assuming additivity. The interaction with degree of urbanization and the study center-specific heterogeneity suggested further variation and unknown factors contributing to diabetes and obesity risk in urban, suburban and rural areas.

### 4.2. Comparison to previous studies

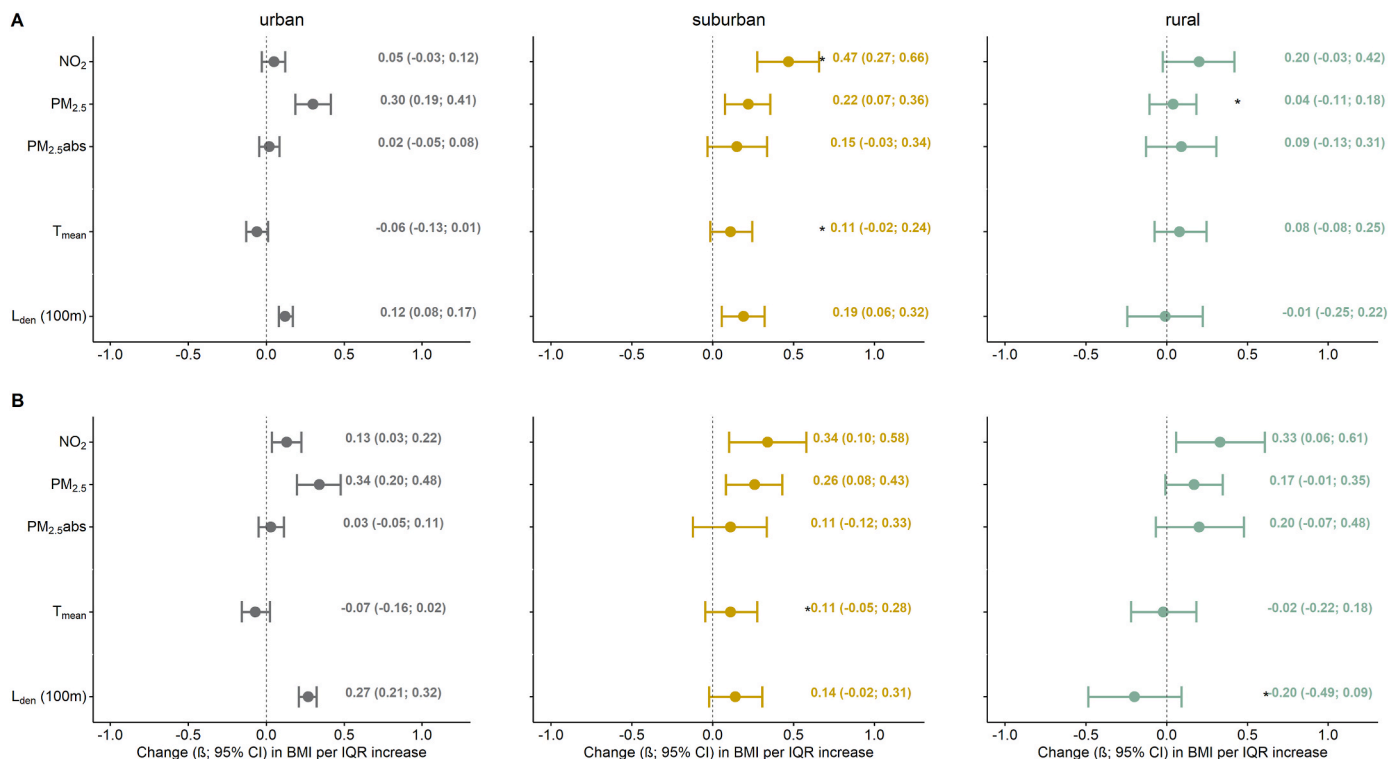
#### 4.2.1. Air pollution

Our findings on the association of adverse levels of  $PM_{2.5}$  with metabolic diseases align with previous studies and analyses. Several meta-analyses found an overall association with PM and diabetes and obesity markers, whereas associations with other air pollutants such as  $NO_2$ ,  $PM_{10}$ , or ozone were mixed (He et al., 2017; Yang et al., 2020a, 2020b; An et al., 2018; Huang et al., 2020). In our study, we show that the associations of  $NO_2$  were reversed in two-exposure models, suggesting that  $PM_{2.5}$  may have confounded the associations. In contrast, studies with higher levels of  $NO_2$  or other air pollutants have observed positive associations with diabetes and obesity-related measures (Pan et al., 2023; Han et al., 2023; Li et al., 2022). This suggest that our findings may not be generalizable to highly pollutant areas. Nevertheless, annual mean residential levels of  $NO_2$  and  $PM_{2.5}$  exceeded the WHO



(caption on next page)

**Fig. 2.** Joint associations of exposure to environmental mixture with diabetes and BMI for n = 86,710 men and n = 88,245 women from the German National Cohort (NAKO). Legend: Joint odds ratios and betas are given per IQR increase across all exposures derived from multi-exposure quantile g-computation models. We used the inverse of NDVI; and bars represent exposure weights, indicating the proportion of each exposure contributing to the overall positive and negative association. Models were additionally adjusted for age, study center region, lifestyle factors, education, unemployment rate at district level and population density. Abbreviations: BMI = Body Mass Index, CI = confidence interval, IQR = interquartile range, L<sub>den</sub> = day-evening-night noise level, NDVI = normalized difference vegetation index, OR = Odds ratio, PM<sub>2.5</sub> = particulate matter with diameter <2.5 μm.



**Fig. 3.** Urbanization-specific associations of environmental exposures with diabetes for men (A) and women (B) in the German National Cohort (NAKO). Legend: Betas and 95%-CI were derived from linear regression models including an interaction term between exposure and degree of urbanization from EUROSTAT. Star indicates  $p_{\text{interaction}} < 0.05$ . All models were additionally adjusted for age, study center region, lifestyle factors, education, unemployment rate and population density. Urban: n = 62,004 men; n = 62,975 women, suburban: n = 13,771 men, n = 14,152 women, rural: n = 10,933 men, n = 11,112 women. Abbreviation: BMI = Body Mass Index, CI = confidence interval, IQR = interquartile range, L<sub>den</sub> = day-evening-night noise level, NO<sub>2</sub> = nitrogen dioxide, OR = odds ratio, PM<sub>2.5</sub> = particulate matter <2.5 μm, PM<sub>2.5abs</sub> = PM<sub>2.5</sub> absorbance, T<sub>mean</sub> = mean air temperature.

recommended levels (WHO, 2021). Reviews by Rajagopalan et al. (Rajagopalan and Brook, 2012; Rajagopalan et al., 2024) gave a summary of the evident mechanistic insights of how air pollutants, in particular PM, affect human metabolism. For example, studies in humans and mice showed a deterioration of endothelial function and insulin sensitivity (Rajagopalan and Brook, 2012; Rajagopalan et al., 2024). Local inflammation in visceral adipose tissue and liver were reported, characterized by increased number of macrophages, oxidative stress, and mitochondrial dysfunction, all linking air pollution to diabetes and obesity (Rajagopalan and Brook, 2012; Rajagopalan et al., 2024).

#### 4.2.2. Air temperature

For annual mean air temperature, our results were mixed and inconclusive, so we could not draw clear conclusions about its effect on diabetes and obesity-related measures. In contrast, two studies using data from the Spanish “Di@bet.es Study” found that an increase from the lowest to the highest quartile of annual air temperature was associated with 1.39 and 1.38 higher odds of diabetes and obesity, respectively (Valdes et al., 2014, 2019). However, Wallwork et al. (2017) and Speakman and Heidari-Bakavoli (2016) found that increases in air temperature were associated with diabetes prevalence, but not with obesity. These studies reported annual mean air temperature ranges of

more than 10 °C, which shows a higher variability than in our study (6 °C). Previous studies on potential mechanistic pathways have focused on the effects of short-term exposure to air temperature, arguing that exposure to cold may increase activity of brown adipose tissue, which modulates triglyceride accumulation and is associated with improved insulin sensitivity (Valdes et al., 2014; Marlatt and Ravussin, 2017). Furthermore, air temperature may also act via lifestyle changes, as optimal temperatures may increase active transportation and physical activity compared to non-optimal temperatures. However, we assumed that air temperature variability in our study was limited and therefore, would not affect physical activity behavior of participants, as confirmed by our sensitivity analysis. Consequently, multi-country studies which provide more air temperature variability may give important insights into the association of air temperature with metabolic health.

#### 4.2.3. Road traffic noise

We identified robust associations of road traffic noise with diabetes and obesity-related measures, independent of other environmental exposures. Furthermore, noise was the main contributor to the multi-exposure associations. This is in line with previous studies that have found consistent associations between noise exposure and several metabolic-related markers (Wang et al., 2020; Gui et al., 2022). In a subsample of the NAKO study, we also demonstrated an association of

road traffic noise exposure with adipose tissue depots and hepatic fat content measured by whole-body magnetic resonance imaging (Niedermayer et al., 2025). There are multiple reviews summarizing the harmful effects of noise exposure on human metabolism (Munzel et al., 2024; Sivakumaran et al., 2022). Noise, especially during night, can disrupt sleep and affects its quality (Smith et al., 2022; Cappuccio and Miller, 2017). This can result in metabolic abnormalities, characterized by an imbalance in hormones regulating appetite and satiety, which can further lead to an imbalance between energy intake and expenditure (Stevens et al., 2023). The other pathway involves an overactive hypothalamic-pituitary-adrenal axis and sympathetic nervous system that comes along with increased release of stress hormones such as adrenaline, cortisol or catecholamines, that in turn leads to increased oxidative stress and a constant state of inflammation (Munzel et al., 2024). We need to note that road traffic noise exposure is only available at urban areas with >100,000 inhabitants and along major roads with more than 3 million vehicles per year (European Commission, 2002). However, we were lacking adequate noise exposure assessment in rural and remote areas, nor were we able to take into account other potential sources of noise, such as aircraft or railway, which may have an additional impact on metabolic health (Bozgar et al., 2024).

#### 4.2.4. Greenness

We observed non-linear associations of greenness with diabetes and obesity-related measures, which is comparable to numerous previous studies, although the described shape of the exposure-response function varied between studies (Dempsey et al., 2018; James et al., 2017; Klompaker et al., 2018). Non-linear associations may account for previous mixed evidence on the effect of surrounding greenness on metabolic health (De la Fuente et al., 2020; Luo et al., 2020). In our study, we observed that these non-linear associations were driven by urban areas, suggesting potential residual confounding. We hypothesize that high NDVI levels in urban areas may be attributed to highly vegetated parks and areas on the outskirts of cities. On the other side, central urban areas may offer more opportunities for indoor physical activity with higher numbers of gyms, which may have resulted in the unexpected association of lower NDVI being related to lower odds of diabetes, obesity and lower BMI and waist circumference. Klompaker et al. (2018) reported that natural greenness is associated with lower odds of overweight and higher physical activity levels, whereas contrary associations were found for urban green. In addition, NDVI as an indicator of greenness has some disadvantages, as it cannot distinguish between different types of greenness, such as parks, grassland, or agricultural areas - and therefore does not reflect the complexity or functional value of green spaces. It also does not provide information on accessibility or actual time spent in green spaces. These limitations may contribute to the observed inverted U-shape, as vegetation density alone may not fully capture the health-relevant characteristics of green spaces (Dempsey et al., 2018; Evenson et al., 2019). These aspects have shown to be important determinants of whether greenness positively affect mental health and physical activity behavior (Dempsey et al., 2018; Villeneuve et al., 2018; Astell-Burt et al., 2014).

#### 4.2.5. Joint associations

Recently, the question of how multiple environmental exposures jointly contribute to disease risk has been raised, but only a few studies have investigated metabolic health outcomes (Klompaker et al., 2019; Sorensen et al., 2022b; Zhang et al., 2024; Chen et al., 2024; Voss et al., 2021). A study by Klompaker et al. (2019) used the cumulative risk index to quantify the cumulative effects of exposures to air pollution, noise and greenness and found a joint odds ratio of 1.13 for diabetes, which was higher than the associations found in single-exposure models. Similarly, a study from Denmark showed a cumulative risk index of 1.12 for diabetes when considering air pollution, greenness and noise together (Sorensen et al., 2022b). Applying a quantile-g computation that can handle multi-collinearity, and allows to explore non-additivity

and non-linearity (Keil et al., 2020), we found joint associations of multiple adverse environmental exposures with diabetes and obesity-related measures. Similarly, Zhang et al. (2024) performed a quantile g-computation in two prospective US-based cohorts of female nurses. They observed a negative joint association between their environmental exposures, including air pollution, nighttime noise, greenness, light at night, air temperature, neighborhood SES, and BMI. However, this protective joint association was predominantly driven by neighborhood SES. After dropping the neighborhood SES variable from the models, the environmental exposure mixture was associated with 0.16 kg/m<sup>2</sup> and 0.06 kg/m<sup>2</sup> higher BMI in the two cohorts (Zhang et al., 2024). Partially consistent with our findings, the positive joint associations were mainly driven by the air pollutants NO<sub>2</sub>, PM<sub>2.5</sub> and nighttime noise (Zhang et al., 2024). In contrast, NO<sub>2</sub> seemed to be confounded in our study in multi-exposure models, which resulted in a contribution of NO<sub>2</sub> to the negative joint associations with the outcome diabetes. In addition, we cannot explain why the contribution of weight differed for the joint association between environmental exposures and diabetes in women. Nevertheless, we need to be careful when interpreting the weights of each exposure, as these methods tend to give more weight to the exposure with the least measurement error (Zhang et al., 2024; Chen et al., 2024). Furthermore, we cannot derive weights from models that consider splines. Therefore, we should note that the weights were derived from models with NDVI as a linear predictor, which may not fully capture the complex exposure-response relationship for greenness in the exposure mixture. More research is therefore needed that examines multiple environmental factors together, investigating potential synergistic and interactive effects with methods that are more flexible than quantile g-computation. Methods such as Bayesian kernel machine regression and gradient boosting may be more appropriate for such complex mixture settings as it was observed in our study (Chen et al., 2024). Unfortunately, these methods are still too computationally intensive to be implemented in large cohort studies such as NAKO. Therefore, we need to strike a balance between the available methods that can assess joint associations. Quantile g-computation shares the simplicity of WQS regression, but relaxes several of its assumptions, including unidirectionality, linearity and the additivity of individual exposure effects (Keil et al., 2020). Consequently, we decided to implement this method in our cross-sectional study. By exploring different exposure mixture sets, we demonstrated the robustness of joint estimates when including or excluding NDVI as linear predictor. However, including NDVI as non-linear predictor led to changes in the effect estimates and larger CI.

#### 4.2.6. Degree of urbanization

We observed that adjusting for population density had a strong impact on the associations of air pollution with obesity-related markers. Population density was included as a confounder, as we hypothesize related to both exposure variation and health outcomes through multiple behavioral and environmental pathways. For example, population density may be a proxy for several infrastructure-related variables such as a higher proportion of active travel options, increased walkability and connectivity, a diverse food environment (Ohanyan et al., 2022; Sun et al., 2022; Sarkar et al., 2017), or explain some population characteristics that differed by sex and SES and that may confound the association of some environmental factors with obesity. With regard to the present study, we hypothesized that study center region and unemployment rate at the district level may have been too coarse to capture these unmeasured confounders. Contributing to this, the associations of air pollution and road traffic noise were most pronounced in urban areas. Furthermore, study centers regions with a high proportion of urban areas tended to show expected associations of air pollution with diabetes and obesity-related measures, whereas study centers regions including a mixture of urban, suburban and rural areas contributed to unexpected protective or null findings. In addition to global trends of higher rates of obesity in rural area (Liu et al., 2021; Non-Communicable

Disease Risk Factor Collaboration, 2017; Lee and Um, 2021), these heterogeneous findings indicate that additional factors may contribute to the development of diabetes and obesity in rural areas, which require further investigation.

#### 4.3. Outlook

Future analyses within NAKO will address several important questions that were beyond the scope of this cross-sectional study. As biomarker data becomes available, studies should investigate potential biological mechanisms through mediation analyses. Ideally, these analyses will use repeated measures of metabolic and inflammatory markers, such as hemoglobin A1c (HbA1c), lipids and C-reactive protein, to better establish temporal order and clarify the pathways linking environmental exposures with diabetes and obesity. Currently, the 10-year follow-up examination within the NAKO cohort is conducted collecting blood samples, anthropometric and clinical biomarkers. In this context, the role of BMI as a potential mediator in the association between environmental exposures and diabetes warrants further investigation in longitudinal studies. As we were primarily interested in the total association rather than the direct association between environmental exposures and diabetes, we did not include BMI in our main adjustment model. However, adjusting for BMI did not alter the results considerably (Table S9). Additionally, genotyping data from the entire cohort will enable the examination of gene–environment interactions, including the application of polygenic risk scores to determine if individuals with a higher genetic predisposition are more susceptible to adverse environmental influences. Once longitudinal follow-up data and genotyping is available, it will be possible to derive more robust causal estimates and calculate population attributable fractions in order to quantify the potential public health impact of reducing environmental exposures.

#### 4.4. Strengths and limitations

Our study has several strengths. First, we used data from a large, deeply phenotyped, multi-center cohort study that provided sufficient exposure contrast, included participants from diverse socio-economic and urbanized environments, and provided several measures of metabolic function. In particular, BMI is critically discussed as measure of obesity (Rubino et al., 2025), but the NAKO study allowed us to compare associations across a range of different measures of adiposity. In addition, we used a novel, powerful method to assess the joint associations of environmental exposure mixtures with diabetes and obesity-related measures. Rather than assessing the association of one exposure while the other exposures are kept constant, the quantile g-computation assesses the associations of the mixture by increasing all exposures by the pre-defined increment (Zhang et al., 2024; Keil et al., 2020). As a result, this method may better reflect reality because it takes into account co-existing exposures and deals with multi-collinearity, thus reducing a potential bias of the association of environmental factors with diabetes and obesity.

We note the following limitations of our study. Firstly, misclassification in diabetes cases may be present as we had to rely on self-reported information on diabetes diagnosis. Although misclassification can bias associations towards or away from the null, our sensitivity analysis indicated a bias towards the null. Excluding participants who did not report the use of anti-diabetic medication tended to strengthen the association, particularly among women. Information on type of diabetes, blood markers, such as hemoglobin A1c, or oral glucose tolerance test results to identify undetected diabetes, was not available at the time of analysis, which would be interesting to investigate in future studies. Secondly, we cannot rule out residual confounding. The proxy of unemployment rate for neighborhood SES was only available at the district level, which was likely too coarse to adequately adjust for potential unmeasured spatial confounders. Thirdly, we did not have information

on diet at the time of the study. However, we assumed that dietary factors might have explained some of the heterogeneity of our participants, but our previous findings in the KORA-Fit study did not show relevant confounding by diet (Niedermayer et al., 2024). In addition, environmental exposures were only available at residential addresses and information on workplace locations or commuting patterns were not available. However, previous studies examining the differences between environmental exposure levels at residential addresses and time-activity-integrated exposure levels have reported strong correlations between the two approaches (often exceeding 0.8), yielding nearly identical health effect estimates (Hoek et al., 2024). Therefore, we assume this will likely bias the associations towards the null. Adding to this, environmental exposures were available at different resolutions and for different exposure time windows. This temporal mismatch may introduce measurement error for different exposures. In our study, we assume that spatial, and for some exposures also temporal, distributions of environmental exposures are relatively stable, as shown in previous studies (Wolf et al., 2025; de Hoogh et al., 2018). Nevertheless, this measurement error for different exposures likely underestimates the true association between environmental exposures and health outcomes. Moreover, it may be a potential reason for null findings, which are hard to explain. Furthermore, measurement error for different exposures may have impacted the assessment of joint associations by quantile g-computation, that these methods tend to give more weight to the exposure with the least measurement error (Zhang et al., 2024; Chen et al., 2024). As mentioned above, weights were derived from models with NDVI as a linear predictor, which may not fully capture the complex exposure-response relationship for greenness in the exposure mixture. Additionally, it should be noted that including NDVI with a spline in the exposure mixture approach resulted in altered effect estimates and larger CIs. Furthermore, we only had cross-sectional data, which did not allow us to observe a temporal sequence of exposure and outcome. Consequently, we only observed associations and not causal effects, which is specifically true for quantile g-computation which aims to assess causal effects. Therefore, future longitudinal studies investigating the effects of multiple environmental exposures jointly on diabetes and obesity are warranted.

#### 5 Conclusions

To conclude, we investigated the associations of six different environmental exposures with diabetes and obesity-related measures in this large multi-center cohort study. We observed that PM<sub>2.5</sub> and road traffic noise were consistently associated with diabetes and measures of obesity, particularly in urban areas, using single-exposure models. NDVI showed an inverted U-shaped exposure-response curve with all outcomes. In addition, joint associations were observed when multiple exposures were considered simultaneously. This suggests that reducing multiple adverse environmental exposures together could be a potential target for preventing diabetes and obesity. Longitudinal studies are needed to corroborate causal effects of co-existing environmental exposures on diabetes and obesity. Moreover, this study highlights the need to identify socio-economic and environmental drivers that can explain differences in diabetes and obesity-related measures between urban and rural areas in future studies.

#### CRedit authorship contribution statement

**Fiona Niedermayer:** Writing – review & editing, Writing – original draft, Visualization, Methodology, Formal analysis, Data curation, Conceptualization. **Barbara Hoffmann:** Writing – review & editing, Supervision, Resources, Methodology. **Boya Zhang:** Writing – review & editing, Methodology. **Jie Chen:** Writing – review & editing, Methodology. **Jaime E. Hart:** Writing – review & editing, Methodology. **Francine Laden:** Writing – review & editing, Methodology. **Gabriele Bolte:** Writing – review & editing, Resources. **Tobia Lakes:** Writing – review & editing, Resources. **Tamara Schikowski:** Writing – review & editing, Resources. **Karin Halina Greiser:** Writing – review & editing,

Resources. **Jeroen Staab**: Writing – review & editing, Resources, Data curation. **Nikolaos Nikolaou**: Writing – review & editing, Resources, Data curation. **Marco Dallavalle**: Writing – review & editing, Resources, Data curation. **Matthias B. Schulze**: Writing – review & editing, Resources. **Wolfgang Lieb**: Writing – review & editing, Resources. **Cara Övermöhle**: Writing – review & editing, Resources. **Thaddäus Tönnies**: Writing – review & editing, Resources. **Verena Katzke**: Writing – review & editing, Resources. **Heiko Becher**: Writing – review & editing, Resources. **Beate Fischer**: Writing – review & editing, Resources. **Michael Leitzmann**: Writing – review & editing, Resources. **Klaus Berger**: Writing – review & editing, Resources. **Fatemeh Mayvaneh**: Writing – review & editing, Resources. **Thomas Keil**: Writing – review & editing, Resources. **Lilian Krist**: Writing – review & editing, Resources. **Carolina J. Klett-Tammen**: Writing – review & editing, Resources. **Jana-Kristin Heise**: Writing – review & editing, Resources. **Tobias Pischon**: Writing – review & editing, Resources. **Ilais Moreno Velásquez**: Writing – review & editing, Resources. **Börge Schmidt**: Writing – review & editing, Resources. **Rajini Nagrani**: Writing – review & editing, Resources. **Stefan Rach**: Writing – review & editing, Resources. **Hermann Brenner**: Writing – review & editing, Resources. **Bernd Hollecsek**: Writing – review & editing, Resources. **Volker Harth**: Writing – review & editing, Resources. **Nadia Obi**: Writing – review & editing, Resources. **Anna Köttgen**: Writing – review & editing, Resources. **Rafał Mikolajczyk**: Writing – review & editing, Resources. **Claudia Meinke-Franze**: Writing – review & editing, Resources. **Wolfgang Hoffmann**: Writing – review & editing, Resources. **Alexandra Schneider**: Writing – review & editing, Resources, Methodology, Data curation. **Kathrin Wolf**: Writing – review & editing, Resources, Methodology, Data curation. **Annette Peters**: Writing – review & editing, Supervision, Resources, Methodology, Funding acquisition, Conceptualization.

## Notes

The authors have no conflicts of interest to declare.

## Ethics approval and consent to participate

The German National Cohort (NAKO) was approved by the initial vote of the ethics committee of the Bavarian Medical Association (“Bayerische Landesärztekammer” (BLÄK), protocol code 13023), followed by votes from all local on-site institutional review boards, and written informed consent of all participants was obtained at the time of study enrollment. The study was conducted in accordance with the Declaration of Helsinki of 1975 (in the current, revised version).

## Data sharing

The datasets analyzed during the current study are not publicly available. Access to and use of NAKO data and biosamples can be obtained via an electronic application portal (<https://transfer.nako.de>). Analysis codes are available from the authors upon request.

## Declaration of generative AI and AI-assisted technologies in the manuscript preparation process

During the preparation of this work the authors used DeepL in order to improve readability and language. After using this tool/service, the authors reviewed and edited the content as needed and take full responsibility for the content of the published article.

## Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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## Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.envres.2026.124096>.

## Data availability

The authors do not have permission to share data.

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